

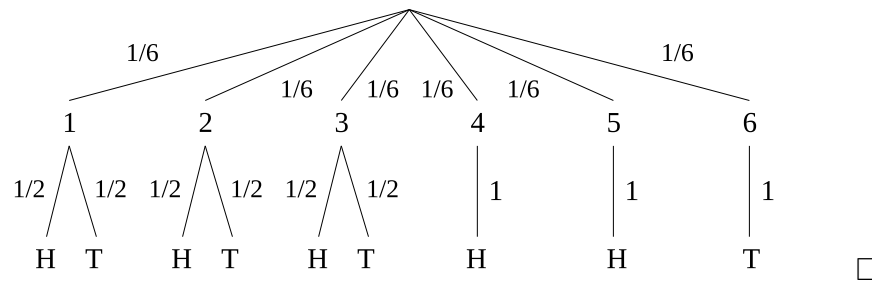
Mathematics 1550H – Introduction to probability
TRENT UNIVERSITY, Winter 2017
Solutions to the Quizzes

Quiz #1. Thursday, 19 January, 2017. [12 minutes]

A standard die is rolled once. If it comes up 1, 2, or 3, you are given a fair coin and toss it once; if the die comes up 4 or 5, you are given a two-headed coin and toss it once; and if it comes up 6, you are given a two-tailed coin and toss it once.

1. Draw the tree diagram for this experiment. [2]
2. What is the sample space Ω for this experiment? [1.5]
3. Determine the probability that a head came up on the coin toss in the second stage of the experiment. [1.5]

SOLUTION. 1. Here is the complete tree diagram for this experiment:



2. $\Omega = \{1H, 1T, 2H, 2T, 3H, 3T, 4H, 5H, 6T\}$ \square

3. If A is the event that a head came up on the coin toss in the second stage of the experiment, then, as a collection of outcomes, $A = \{1H, 2H, 3H, 4H, 5H\}$. Thus, getting the probabilities of the individual outcomes from the tree diagram given in the solution to question 1 above, we have:

$$\begin{aligned}
 P(A) &= m(1H) + m(2H) + m(3H) + m(4H) + m(5H) \\
 &= \frac{1}{6} \cdot \frac{1}{2} + \frac{1}{6} \cdot \frac{1}{2} + \frac{1}{6} \cdot \frac{1}{2} + \frac{1}{6} \cdot 1 + \frac{1}{6} \cdot 1 \\
 &= \frac{1}{12} + \frac{1}{12} + \frac{1}{12} + \frac{1}{6} + \frac{1}{6} = \frac{1}{12} + \frac{1}{12} + \frac{1}{12} + \frac{2}{12} + \frac{2}{12} = \frac{7}{12} \blacksquare
 \end{aligned}$$

Quiz #2. Thursday, 26 January, 2017. [10 minutes]

A hand of five cards is drawn simultaneously (and randomly!) from a standard 52-card deck.

1. What is the probability that the hand has exactly two cards of the same kind? [5]

SOLUTION. First, there are $\binom{52}{5} = \frac{52!}{(52-5)!5!} = \frac{52 \cdot 51 \cdot 50 \cdot 49 \cdot 48}{5 \cdot 4 \cdot 3 \cdot 2 \cdot 1} = \frac{311875200}{120} = 2598960$ ways to pick out 5 cards out of 52 cards if order doesn't matter. Note that each of these hands is as likely as any other to be drawn.

Second, there are $\binom{13}{1} = 13$ ways to choose the kind for the pair and $\binom{4}{2} = 6$ ways to choose a pair from the four cards of that kind in the deck. Similarly, there are $\binom{12}{3} = \frac{12!}{(12-3)!3!} = \frac{12 \cdot 11 \cdot 10}{3 \cdot 2 \cdot 1} = \frac{1320}{6} = 220$ ways to pick three more kinds for the remaining cards in the hand, and, for each of those three kinds, $\binom{4}{1} = 4$ ways to pick a card of that kind. It follows that there are

$$\binom{13}{1} \binom{4}{2} \binom{12}{3} \binom{4}{1} \binom{4}{1} \binom{4}{1} = 13 \cdot 6 \cdot 220 \cdot 4 \cdot 4 \cdot 4 = 1098240$$

possible hands with exactly two cards of the same kind.

Finally, since each hand is as likely to be drawn as any other, it follows that:

$$P(\text{exactly 2 cards of the same kind}) = \frac{\binom{13}{1} \binom{4}{2} \binom{12}{3} \binom{4}{1} \binom{4}{1} \binom{4}{1}}{\binom{52}{5}} = \frac{1098240}{2598960} \approx 0.42$$

An unsimplified answer is just fine for this quiz, but in the real world you would probably want the decimal, quite possibly to more places than given here. ■

Quiz #3. Thursday, 2 February, 2017. [10 minutes]

A fair coin is tossed four times. Let A be the event that the second toss was a tail and let B be the event that there were exactly two tails among the four tosses.

1. Compute $P(A|B)$, the conditional probability of A given B . [4]
2. Determine whether the events A and B are independent. [1]

SOLUTION. 1. We need to compute $P(B)$ and $P(A \cap B)$ in order to compute $P(A|B) = \frac{P(A \cap B)}{P(B)}$. Note that if a fair coin is tossed four times, there are $2^4 = 16$ equally likely outcomes.

First, there are $\binom{4}{2} = 6$ outcomes in B , as this is the number of ways one could choose exactly two of the four tosses to be those in which a tail occurred. This means that $P(B) = \frac{6}{16}$.

Second, $A \cap B$ consists of those outcomes in which a tail occurred on the second toss and exactly one tail occurred among the other three tosses. There are $\binom{3}{1} = 3$ such outcomes as there are three other tosses from which to choose the one other toss in which a tail could occur. This means that $P(A \cap B) = \frac{3}{16}$.

It follows that $P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{3/16}{6/16} = \frac{3}{6} = \frac{1}{2}$. □

2. Since the coin is fair, the probability that it comes up tails on any particular toss is $\frac{1}{2}$, so, in particular, $P(A) = \frac{1}{2}$. Since $P(A|B) = \frac{1}{2} = P(A)$, the events A and B are independent. ■

Quiz #4. Thursday, 9 February, 2017. [10 minutes]

A fair standard six-sided die is rolled four times.

1. What is the expected value of the number of rolls in which the die comes up with a multiple of three? [5]

SOLUTION. When the fair die is rolled once, each of the six faces is as likely to come up as any other, so each has a probability of $\frac{1}{6}$ of coming up. The multiples of three among the integers 1 through 6 are 3 and 6, so it follows that

$$P(\text{multiple of 3}) = m(3) + m(6) = \frac{1}{6} + \frac{1}{6} = \frac{2}{6} = \frac{1}{3}.$$

Rolling the die four times and counting the rolls in which a multiple of 3 came up is like repeating a Bernoulli trial with a probability of success of $\frac{1}{3}$ four times. This has a binomial distribution with $n = 4$, $p = \frac{1}{3}$, and $q = 1 - \frac{1}{3} = \frac{2}{3}$, so it has an expected value of $E(X) = np = 4 \cdot \frac{1}{3} = \frac{4}{3}$. ■

Quiz #5. Thursday, 16 February, 2017. [12 minutes]

The blades of grass on a certain lawn are all between 3.5 and 6.5 *cm* long. The random variable X is the length of randomly chosen blade of grass from this lawn, and is as likely to have any given value between 3.5 and 6.5 *cm* as it is to have any other value between these extremes.

1. What is the probability density function of X ? [2]
2. Verify that the function you gave to answer question 1 is a valid probability density. [2]
3. Find the expected value, $E(X)$, of X . [1]

SOLUTION. 1. Denote the probability density function of X by $f(x)$. Since all the values that X takes on are between 3.5 and 6.5, we must have $f(x) = 0$ when $x < 3.5$ or $x > 6.5$; since we also have that any value between 3.5 and 6.5 is equally likely, X has a continuous uniform distribution on $[3.5, 6.5]$, so for $3.5 \leq x \leq 6.5$ we have $f(x) = \frac{1}{6.5 - 3.5} = \frac{1}{3}$. Thus the probability density function of X is $f(x) = \begin{cases} \frac{1}{3} & 3.5 \leq x \leq 6.5 \\ 0 & x < 3.5 \text{ or } x > 6.5 \end{cases}$. □

2. First, $f(x) \geq 0$ for all x by its definition: $f(x) = \frac{1}{3} > 0$ for $3.5 \leq x \leq 6.5$, and $f(x) = 0$ otherwise. Second,

$$\begin{aligned} \int_{-\infty}^{\infty} f(x) dx &= \int_{-\infty}^{3.5} 0 dx + \int_{3.5}^{6.5} \frac{1}{3} dx + \int_{6.5}^{\infty} 0 dx = 0 + \frac{1}{3}x \Big|_{3.5}^{6.5} + 0 \\ &= \frac{1}{3} \cdot 6.5 - \frac{1}{3} \cdot 3.5 = \frac{6.5 - 3.5}{3} = \frac{3}{3} = 1, \end{aligned}$$

so $f(x)$ meets both requirements for being a valid probability density. □

3. As noted in class, the expected value of a continuous random variable with a uniform distribution on an interval $[a, b]$ has expected value $\frac{a+b}{2}$. Since, as observed in the solution to question 1 above, the given random variable X has uniform distribution on $[3.5, 6.5]$, its expected value is $E(X) = \frac{3.5 + 6.5}{2} = \frac{10}{2} = 5$. ■

Quiz #6. Thursday, 9 March, 2017. [10 minutes]

A fair four-sided die has its faces labelled with the numbers -1 , -2 , $+3$, and $+4$. It is rolled once and the random variable X returns the number on the face that comes up.

1. Find the probability function of X . [1]
2. Compute the expected value $E(X)$ of X . [2]
3. Compute the variance $V(X)$ of X . [2]

SOLUTION. 1. Since the die is fair, each of its four faces is as likely to come up as any other, giving X a discrete uniform distribution, *i.e.* $m(-1) = m(-2) = m(+3) = m(+4) = \frac{1}{4} = 0.25$. \square

2. By definition:

$$E(X) = \sum_x xm(x) = (-1) \cdot \frac{1}{4} + (-2) \cdot \frac{1}{4} + 3 \cdot \frac{1}{4} + 4 \cdot \frac{1}{4} = \frac{-1 - 2 + 3 + 4}{4} = \frac{4}{4} = 1 \quad \square$$

3. We will use the formula $V(X) = E(X^2) - [E(X)]^2$, so we first need to compute $E(X^2)$:

$$E(X^2) = \sum_x x^2 m(x) = (-1)^2 \cdot \frac{1}{4} + (-2)^2 \cdot \frac{1}{4} + 3^2 \cdot \frac{1}{4} + 4^2 \cdot \frac{1}{4} = \frac{1 + 4 + 9 + 16}{4} = \frac{30}{4} = \frac{15}{2}$$

It follows that $V(X) = E(X^2) - [E(X)]^2 = \frac{15}{2} - 1 = \frac{15}{2} - \frac{2}{2} = \frac{13}{2} = 6.5$. \blacksquare

Quiz #7. ~~Thursday, 16~~ Tuesday, 21 March, 2017. [10 minutes]

Answer *one* (1) of the following questions.

1. Show that $E(X^2) \geq [E(X)]^2$ for any random variable X . [5]
2. Suppose X_1, X_2, \dots, X_9 are independent and identically distributed random variables, with $E(X_i) = 41$ and variance $V(X_i) = 81$ for each i . Explain as fully as you can why $\frac{X_1 + X_2 + \dots + X_9}{9}$ is likely to be closer to 41 than X_1 is by itself. [5]

SOLUTION. 1. The very short, but good enough, version is that $V(X) = E(X^2) - [E(X)]^2$ and – as noted in class – is always positive, which is only possible if $E(X^2) \geq [E(X)]^2$.

The complete to the point of paranoia version is that $[X - E(X)]^2 \geq 0$ because it is a square, from which it follows that $V(X) = E([X - E(X)]^2) \geq 0$ too. The computation given in class, exploiting the linearity of expected value and the fact that $E(X)$ is a constant,

$$\begin{aligned} V(X) &= E([X - E(X)]^2) = E(X^2 - 2E(X) \cdot X + [E(X)]^2) \\ &= E(X^2) - E(2E(X) \cdot X) + E([E(X)]^2) = E(X^2) - 2E(X) \cdot E(X) + [E(X)]^2 \\ &= E(X^2) - 2[E(X)]^2 + [E(X)]^2 = E(X^2) - [E(X)]^2, \end{aligned}$$

then implies that $E(X^2) - [E(X)]^2 \geq 0$, so $E(X^2) \geq [E(X)]^2$. \square

2. As noted in class, $\frac{X_1 + X_2 + \dots + X_9}{9}$ has the same expected value as any of the identically distributed X_i . Observe that because the X_i are independent, the properties of variance tell us that:

$$\begin{aligned} V\left(\frac{X_1 + X_2 + \dots + X_9}{9}\right) &= \left(\frac{1}{9}\right) V(X_1 + X_2 + \dots + X_9) \\ &= \frac{1}{81} [V(X_1) + V(X_2) + \dots + V(X_9)] \\ &= \frac{1}{81} [81 + 81 + \dots + 81] = \frac{1}{81} [9 \cdot 81] = 9 \end{aligned}$$

Thus $V\left(\frac{X_1 + X_2 + \dots + X_9}{9}\right) = 9 < 81 = V(X_1)$, and since variance is a measure of how close to its expected value a random variable is likely to end up, $\frac{X_1 + X_2 + \dots + X_9}{9}$ is likely to be closer to 41 than X_1 is by itself. \blacksquare

Quiz #8. Thursday, 23 March, 2017. [10 minutes]

1. How many distinguishable ways are there to take all the letters, including the repetitions, in the word “Peterborough” and arrange them in a row? [5]

SOLUTION. “Peterborough” has twelve letters, counting the repetitions: there are two “e”s, two “o”s, and two “r”s. There are $12!$ ways to arrange twelve distinct things, but for each of the three identical pairs of letters there are $2!$ ways to rearrange them among themselves without the change being apparent. It follows that there are $\frac{12!}{2!2!2!} = 59875200$ distinguishable ways to arrange all the letters in the word “Peterborough”. ■

Quiz #9. Thursday, 30 March, 2017. [10 minutes]

Suppose X is a random variable with $\mu = E(X) = 1$ and $\sigma = \sqrt{V(X)} = 2$.

1. What can you deduce about $P(|X - 1| \geq 4)$ using Chebyshev’s Inequality if you know nothing else about X ? [2]
2. What can you deduce about $P(X \geq 5)$ using Chebyshev’s Inequality if you also know that X is symmetric about $\mu = 1$ (i.e. $P(X - 1 \geq a) = P(X - 1 \leq -a)$ for all a)? [1.5]
3. What can you deduce about $P(X \geq 5)$ using Chebyshev’s Inequality if instead you know that X is never negative? [1.5]

SOLUTION. 1. Chebyshev’s Inequality states that $P(|X - \mu| \geq t) \leq \frac{\sigma^2}{t^2}$. We have $\mu = 1$ and $\sigma = 2$, so looking at $P(|X - 1| \geq 4)$ we see that we must have $t = 4$. Thus $P(|X - 1| \geq 4) \leq \frac{2^2}{4^2} \leq \frac{4}{16} = \frac{1}{4} = 0.25$. □

2. First, note that $P(X \geq 5) = P(X - 1 \geq 4)$. Second, because X is symmetric about $\mu = 1$, we have $P(|X - 1| \geq 4) = P(X - 1 \geq 4) + P(X - 1 \leq -4) = 2P(X - 1 \geq 4)$. Using the answer to question 1 above, this means that $P(X - 1 \geq 4) = \frac{1}{2}P(|X - 1| \geq 4) \leq \frac{1}{2} \cdot \frac{1}{4} = \frac{1}{8}$. □

3. If X is never negative, then $|X - 1| \geq 4$ only when $X - 1 \geq 4$, i.e. when $X \geq 5$. Thus $P(X \geq 5) = P(X - 1 \geq 4) = P(|X - 1| \geq 4)$, which, by Chebyshev’s Inequality, is $\leq \frac{1}{4} = 0.25$, as in the answer to question 1 above. That is, we get the same answer as in question 1 in this case ... □

Quiz #10. Thursday, 6 April, 2017. [10 minutes]

Suppose the discrete random variables X and Y are jointly distributed according to the following table:

$x \backslash Y$	1	2
1	0.1	0.2
2	0.2	0.2
3	0.1	0.2

1. Compute the expected values $E(X)$ and $E(Y)$, the variances $V(X)$ and $V(Y)$, and the covariance $\text{Cov}(X, Y)$ of X and Y . [5]

SOLUTION. Here we go:

$$E(X) = 1 \cdot P(X = 1) + 2 \cdot P(X = 2) + 3 \cdot P(X = 3)$$

$$= 1(0.1 + 0.2) + 2(0.2 + 0.2) + 3(0.1 + 0.2) = 0.3 + 0.8 + 0.9 = 2$$

$$E(Y) = 1 \cdot P(Y = 1) + 2 \cdot P(Y = 2) = 1(0.1 + 0.2 + 0.1) + 2(0.2 + 0.2 + 0.2)$$

$$= 0.4 + 1.2 = 1.6$$

$$E(X^2) = 1^2(0.1 + 0.2) + 2^2(0.2 + 0.2) + 3^2(0.1 + 0.2) = 1 \cdot 0.3 + 4 \cdot 0.4 + 9 \cdot 0.3$$

$$= 0.3 + 1.6 + 2.7 = 4.6$$

$$V(X) = E(X^2) - [E(X)]^2 = 4.6 - 2^2 = 4.6 - 4 = 0.6$$

$$E(Y^2) = 1^2(0.1 + 0.2 + 0.1) + 2^2(0.2 + 0.2 + 0.2) = 1 \cdot 0.4 + 4 \cdot 0.6$$

$$= 0.4 + 2.4 = 2.8$$

$$V(Y) = E(Y^2) - [E(Y)]^2 = 2.8 - (1.6)^2 = 2.8 - 2.56 = 0.24$$

$$E(XY) = 1 \cdot 1 \cdot P(X = 1 \& Y = 1) + \dots + 3 \cdot 2 \cdot \dots \cdot P(X = 3 \& Y = 2)$$

$$= 1 \cdot 1 \cdot 0.1 + 1 \cdot 2 \cdot 0.2 + 2 \cdot 1 \cdot 0.2 + 2 \cdot 2 \cdot 0.2 + 3 \cdot 1 \cdot 0.1 + 3 \cdot 2 \cdot 0.2$$

$$= 0.1 + 0.4 + 0.4 + 0.8 + 0.3 + 1.2 = 3.2$$

$$\text{Cov}(X, Y) = E(XY) - E(X) \cdot E(Y) = 3.2 - 2 \cdot 1.6 = 3.2 - 3.2 = 0$$

Whew! ■